# The Weibull Distribution The Mathematics 

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In economics we can use the Weibull distribution to model (1) random event arrival time and (2) random number of events realized within a given time interval. In this white paper we will develop the mathematics for the Weibull distribution. To that end we will work through the following hypothetical problem...

## Our Hypothetical Problem

We are tasked with answering the following questions applicable to a Weibull distribution with shape parameter = 1.50 and scale parameter $=2.75$.

Question 1: Graph the probability distribution where $\kappa$ (shape) $=1.50$ and $\lambda$ (scale) $=2.75$.
Question 2: What is the mean and variance of the distribution?
Question 3: What is the probability that the event will arrive after 3.00 years?
Question 4: Given that we pulled 0.35 from a uniform distribution what is the event arrival time?

## The Weibull Distribution

We will define the variable $t$ to be time in years and the function $f(t)$ to be the probability density function (PDF) of a Weibull-distributed random variable $t$. The equation for the probability density function where the variable $\kappa$ is the shape parameter and the variable $\lambda$ is the scale parameter is...

$$
\begin{equation*}
\text { PDF : } f(t)=\frac{\kappa}{\lambda}\left(\frac{t}{\lambda}\right)^{\kappa-1} \operatorname{Exp}\left\{-\left(\frac{t}{\lambda}\right)^{\kappa}\right\} \ldots \text { where... } t \geq 0, \kappa>0, \lambda>0 \tag{1}
\end{equation*}
$$

Note that when the shape parameter $\kappa<1$ the probability density function $f(t)$ at $t=0$ is undefined. Using Equation (1) above this statement in equation form is...

$$
\begin{equation*}
\left.f(0)=\frac{\kappa}{\lambda}\left(\frac{0}{\lambda}\right)^{\kappa-1} \operatorname{Exp}\left\{-\left(\frac{0}{\lambda}\right)^{\kappa}\right\}=\frac{\kappa}{\lambda}(0)^{\kappa-1} \right\rvert\, f(0)=0 . . \text { when... } \kappa \geq 1 \mid f(0)=\infty . . \text { when... } \kappa<1 \tag{2}
\end{equation*}
$$

Note that when the shape parameter $\kappa=1$ the Weibull distribution PDF becomes the Exponential distribution PDF. Using Equation (1) above this statement in equation form is... [1]

$$
\begin{equation*}
f(t)=\frac{1}{\lambda}\left(\frac{t}{\lambda}\right)^{1-1} \operatorname{Exp}\left\{-\left(\frac{t}{\lambda}\right)^{1}\right\}=a \operatorname{Exp}\{-a t\} \ldots \text { where... } a=\frac{1}{\lambda} \tag{3}
\end{equation*}
$$

Using Equation (1) above the equation for the cumulative probability that the Weibull-distributed random variable $t$ is greater than or equal to $a$ and less than or equal to $b$ is...

$$
\begin{equation*}
\operatorname{Prob}[a \leq t \leq b]=\int_{a}^{b} f(t) \delta t=\int_{a}^{b} \frac{\kappa}{\lambda}\left(\frac{t}{\lambda}\right)^{\kappa-1} \operatorname{Exp}\left\{-\left(\frac{t}{\lambda}\right)^{\kappa}\right\} \delta t \tag{4}
\end{equation*}
$$

Using Appendix Equation (23) below the solution to Equation (4) above is...

$$
\begin{equation*}
\operatorname{Prob}[a \leq t \leq b]=\operatorname{Exp}\left\{-\left(\frac{a}{\lambda}\right)^{k}\right\}-\operatorname{Exp}\left\{-\left(\frac{b}{\lambda}\right)^{k}\right\} \tag{5}
\end{equation*}
$$

The cumulative distribution function (CDF) is the probability that the random variable $t$ drawn from a Weibull distribution will take a value less than or equal to z. Using Appendix Equation (24) below the equation for the cumulative distribution function is...

$$
\begin{equation*}
\mathrm{CDF}: \operatorname{Prob}[t \leq z]=\int_{0}^{z} f(t) \delta t=1-\operatorname{Exp}\left\{-\left(\frac{z}{\lambda}\right)^{k}\right\} \ldots \text { where... } t \geq 0, \kappa>0, \lambda>0 \tag{6}
\end{equation*}
$$

## Distribution Mean and Variance

Using Equation (1) above the equation for the first moment (FM) of the Weibull-distributed random variable $x$ is...

$$
\begin{equation*}
F M=\int_{0}^{\infty} t f(t) \delta t=\int_{0}^{\infty} t \frac{\kappa}{\lambda}\left(\frac{t}{\lambda}\right)^{\kappa-1} \operatorname{Exp}\left\{-\left(\frac{t}{\lambda}\right)^{\kappa}\right\} \delta t \tag{7}
\end{equation*}
$$

Using Appendix Equations (30) and (31) below the solution to Equation (7) above is...

$$
\begin{equation*}
F M=\lambda \Gamma\left(1+\kappa^{-1}\right)=\lambda \operatorname{EXP}\left(\operatorname{GAMMALN}\left(1+\kappa^{-1}\right)\right) \tag{8}
\end{equation*}
$$

Using Equation (1) above the equation for the second moment (SM) of the Weibull-distributed random variable $x$ is...

$$
\begin{equation*}
S M=\int_{0}^{\infty} t^{2} f(t) \delta t=\int_{0}^{\infty} t^{2} \frac{\kappa}{\lambda}\left(\frac{t}{\lambda}\right)^{\kappa-1} \operatorname{Exp}\left\{-\left(\frac{t}{\lambda}\right)^{\kappa}\right\} \delta t \tag{9}
\end{equation*}
$$

Using Appendix Equations (35) and (36) below the solution to Equation (9) above is...

$$
\begin{equation*}
S M=\lambda^{2} \Gamma\left(1+2 \kappa^{-1}\right)=\lambda^{2} \operatorname{EXP}\left(\operatorname{GAMMALN}\left(1+2 \kappa^{-1}\right)\right) \tag{10}
\end{equation*}
$$

Using Equation (8) above the equation for the mean of a Weibull-distributed random variable $t$ is...

$$
\begin{equation*}
\text { Mean of random variable } t=F M=\lambda \Gamma\left(1+\kappa^{-1}\right) \tag{11}
\end{equation*}
$$

Using Equations (8) and (10) above the equation for the variance of a Weibull-distributed random variable $t$ is...

$$
\begin{equation*}
\text { Variance of random variable } t=S M-F M^{2}=\lambda^{2}\left[\Gamma\left(1+2 \kappa^{-1}\right)-\left(\Gamma\left(1+\kappa^{-1}\right)\right)^{2}\right] \tag{12}
\end{equation*}
$$

## Monte Carlo Simulation

We will define the variable $p$ to be the probability of a given event. Since the value of $p$ lies in the range $[0,1]$ we can pull the random value of $p$ from a uniform distribution with domain $[0,1]$. This statement in equation form is...

$$
\begin{equation*}
p=\text { Random number (i.e. cumulative probability) drawn from a uniform distribution with domain }[0,1] \tag{13}
\end{equation*}
$$

We will define the variable $z$ to be a random variable drawn from a Weibull distribution. Using Equation (6) above we can set the value of $p$ from Equation (13) above to the following equality...

$$
\begin{equation*}
p=1-\operatorname{Exp}\left\{-\left(\frac{z}{\lambda}\right)^{k}\right\} \tag{14}
\end{equation*}
$$

Solving Equation (14) above for $z$ we get the Weibull-distributed random variable $z$ that is equivalent to the random variable $p$ drawn from a uniform distribution...

$$
\begin{equation*}
z=\lambda[-\ln (1-p)]^{\frac{1}{k}} \tag{15}
\end{equation*}
$$

## The Answers To Our Hypothetical Problem

Question 1: Graph the probability distribution where $\kappa$ (shape) $=1.50$ and $\lambda$ (scale) $=2.75$.


Question 2: What is the mean and variance of the distribution?
Using Equations (16) above and (36) below the mean of the distribution is...

$$
\begin{equation*}
\text { Mean }=2.75 \times \Gamma\left(1+1.50^{-1}\right)=2.4825 \tag{16}
\end{equation*}
$$

Using Equations (12) above and (36) below the mean of the distribution is...

$$
\begin{equation*}
\text { Variance }=2.75^{2} \times\left[\Gamma\left(1+2 \times 1.50^{-1}\right)-\left(\Gamma\left(1+1.50^{-1}\right)\right)^{2}\right]=2.8412 \tag{17}
\end{equation*}
$$

Question 3: What is the probability that the event will arrive after 3.00 years?
Using Equation (6) above the answer to the question is...

$$
\begin{equation*}
\operatorname{Prob}[t \geq 3.00]=1-\left(1-\operatorname{Exp}\left\{-\left(\frac{3.00}{2.75}\right)^{1.50}\right\}\right)=1-0.68=0.32 \tag{18}
\end{equation*}
$$

Question 4: Given that we pulled 0.35 from a uniform distribution what is the event arrival time?
Using Equation (15) above the answer to the question is...

$$
\begin{equation*}
z=2.75 \times[-\ln (1-0.35)]^{\frac{1}{1.50}}=1.57 \text { years } \tag{19}
\end{equation*}
$$

## References

[1] Gary Schurman, The Exponential Distribution, March, 2012.
[2] Gary Schurman, The Gamma Function, April, 2016.

## Appendix

A. We will define the functions $f(t)$ and $g(t)$ to be...

$$
\begin{equation*}
f(t)=\frac{\kappa}{\lambda}\left(\frac{t}{\lambda}\right)^{\kappa-1} \operatorname{Exp}\left\{-\left(\frac{t}{\lambda}\right)^{\kappa}\right\} \ldots \text { and... } g(t)=-\operatorname{Exp}\left\{-\left(\frac{t}{\lambda}\right)^{k}\right\} \tag{20}
\end{equation*}
$$

Using Equation (20) above the equation for the derivative of the function $g(t)$ is...

$$
\begin{equation*}
\frac{\delta}{\delta t} g(t)=-\left[-\kappa\left(\frac{t}{\lambda}\right)^{k-1} \times \frac{1}{\lambda} \times \operatorname{Exp}\left\{-\left(\frac{t}{\lambda}\right)^{k}\right\}\right]=\frac{\kappa}{\lambda}\left(\frac{t}{\lambda}\right)^{\kappa-1} \operatorname{Exp}\left\{-\left(\frac{t}{\lambda}\right)^{\kappa}\right\}=f(t) \tag{21}
\end{equation*}
$$

Since the derivative of $g(t)=f(t)$ then the antiderivative of the function $f(t)$ is the function $g(t)$.
B. Using Equations (20) and (21) above the solution to the following integral is...

$$
\begin{equation*}
I=\int_{a}^{b} f(t) \delta t=\int_{a}^{b} \frac{\kappa}{\lambda}\left(\frac{t}{\lambda}\right)^{\kappa-1} \operatorname{Exp}\left\{-\left(\frac{t}{\lambda}\right)^{\kappa}\right\} \delta t=-\operatorname{Exp}\left\{-\left(\frac{t}{\lambda}\right)^{k}\right\}\left[\left[_{t=a}^{t=b}\right.\right. \tag{22}
\end{equation*}
$$

The solution to Equation (22) above is...

$$
\begin{equation*}
I=-\operatorname{Exp}\left\{-\left(\frac{b}{\lambda}\right)^{k}\right\}--\operatorname{Exp}\left\{-\left(\frac{a}{\lambda}\right)^{k}\right\}=\operatorname{Exp}\left\{-\left(\frac{a}{\lambda}\right)^{k}\right\}-\operatorname{Exp}\left\{-\left(\frac{b}{\lambda}\right)^{k}\right\} \tag{23}
\end{equation*}
$$

Note that when the lower and upper bounds of integration in Equation (22) above change to 0 and $z$, respectively, then Equation (23) above becomes...

$$
\begin{equation*}
I=\int_{0}^{z} f(x) \delta x=\operatorname{Exp}\left\{-\left(\frac{0}{\lambda}\right)^{k}\right\}-\operatorname{Exp}\left\{-\left(\frac{z}{\lambda}\right)^{k}\right\}=1-\operatorname{Exp}\left\{-\left(\frac{z}{\lambda}\right)^{k}\right\} \tag{24}
\end{equation*}
$$

Note that Equation (24) above is the cumulative distribution function for the Weibull-distributed random variable $z$.
C. The solution to fhe following integral is...

$$
\begin{equation*}
I=\int_{0}^{\infty} t f(t) \delta t=\int_{0}^{\infty} t \frac{\kappa}{\lambda}\left(\frac{t}{\lambda}\right)^{\kappa-1} \operatorname{Exp}\left\{-\left(\frac{t}{\lambda}\right)^{\kappa}\right\} \delta t \tag{25}
\end{equation*}
$$

We will define the variable $u$ to be...

$$
\begin{equation*}
u=\left(\frac{t}{\lambda}\right)^{\kappa} \ldots \text { such that... } t=\lambda u^{\frac{1}{\kappa}} \ldots \text { and... } t^{2}=\lambda^{2} u^{\frac{2}{\kappa}} \tag{26}
\end{equation*}
$$

The derivative of the variable $u$ with respect to the random variable $t$ is...

$$
\begin{equation*}
\frac{\delta u}{\delta t}=\frac{\kappa}{\lambda}\left(\frac{t}{\lambda}\right)^{\kappa-1} \ldots \text { such that... } \delta u=\frac{\kappa}{\lambda}\left(\frac{t}{\lambda}\right)^{\kappa-1} \delta t \tag{27}
\end{equation*}
$$

Using Equations (26) and (27) above we can rewrite Equation (25) above as...

$$
\begin{equation*}
I=\lambda \int_{0}^{\infty} u^{\frac{1}{\kappa}} \operatorname{Exp}\{-u\} \delta u \tag{28}
\end{equation*}
$$

Note that we can rewrite Equation (28) above as...

$$
\begin{equation*}
I=\lambda \int_{0}^{\infty} u^{\alpha-1} \operatorname{Exp}\{-u\} \delta u \ldots \text { where... } \alpha=1+\frac{1}{\kappa} \tag{29}
\end{equation*}
$$

Noting that the integral in Equation (29) above is the Gamma Function, the solution to that integral is... [2]

$$
\begin{equation*}
I=\lambda \Gamma(\alpha)=\lambda \Gamma\left(1+\kappa^{-1}\right) \tag{30}
\end{equation*}
$$

In Excel the gamma function of $\alpha$ is represented by the following Excel function... [2]

$$
\begin{equation*}
\Gamma(\alpha)=\operatorname{EXP}(\operatorname{GAMMALN}(\alpha)) \ldots \text { where } \ldots \alpha>0 \tag{31}
\end{equation*}
$$

D. The solution to fhe following integral is...

$$
\begin{equation*}
I=\int_{0}^{\infty} t^{2} f(t) \delta t=\int_{0}^{\infty} t^{2} \frac{\kappa}{\lambda}\left(\frac{t}{\lambda}\right)^{\kappa-1} \operatorname{Exp}\left\{-\left(\frac{t}{\lambda}\right)^{\kappa}\right\} \delta t \tag{32}
\end{equation*}
$$

Using Equations (26) and (27) above we can rewrite Equation (32) above as...

$$
\begin{equation*}
I=\lambda^{2} \int_{0}^{\infty} u^{\frac{2}{\kappa}} \operatorname{Exp}\{-u\} \delta u \tag{33}
\end{equation*}
$$

Note that we can rewrite Equation (33) above as...

$$
\begin{equation*}
I=\lambda^{2} \int_{0}^{\infty} u^{\alpha-1} \operatorname{Exp}\{-u\} \delta u \ldots \text { where... } \alpha=1+\frac{2}{\kappa} \tag{34}
\end{equation*}
$$

Noting that the integral in Equation (34) above is the Gamma Function, the solution to that integral is... [2]

$$
\begin{equation*}
I=\lambda^{2} \Gamma(\alpha)=\lambda^{2} \Gamma\left(1+2 \kappa^{-1}\right) \tag{35}
\end{equation*}
$$

In Excel the gamma function of $\alpha$ is represented by the following Excel function... [2]

$$
\begin{equation*}
\Gamma(\alpha)=\operatorname{EXP}(\operatorname{GAMMALN}(\alpha)) \ldots \text { where } \ldots \alpha>0 \tag{36}
\end{equation*}
$$

E. Given the cumulative probability $p$ and solving for the random variable $z \ldots$

$$
\begin{align*}
p & =1-\operatorname{Exp}\left\{-\left(\frac{z}{\lambda}\right)^{k}\right\} \\
\operatorname{Exp}\left\{-\left(\frac{z}{\lambda}\right)^{k}\right\} & =1-p \\
-\left(\frac{z}{\lambda}\right)^{k} & =\ln (1-p) \\
\left(\frac{z}{\lambda}\right)^{k} & =-\ln (1-p) \\
\frac{z}{\lambda} & =[-\ln (1-p)]^{\frac{1}{k}} \\
z & =\lambda[-\ln (1-p)]^{\frac{1}{k}} \tag{37}
\end{align*}
$$

